

Improving Extreme Anchor Tension Prediction of a 10-MW Floating Semi-Submersible Type Wind Turbine, Using Highly Correlated Surge Motion Record

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Gaidai O, Xing Y, Wang F, Wang S, Yan P and Naess A (2022) Improving Extreme Anchor Tension Prediction of a 10-MW Floating Semi-Submersible Type Wind Turbine, Using Highly Correlated Surge Motion Record. Front. Mech. Eng 8:888497. doi: 10.3389/fmech.2022.888497 Extreme value prediction of the load-effect responses of complex offshore structures such as the floating wind turbine (FWT) is crucial in ultimate limit state (ULS) design. This paper considers two cases to understand the feasibility of the bivariate correction on the extreme load and motion responses of a 10-MW semi-submersible type FWT. The empirical anchor tension force and surge motion used in this study are obtained from the FAST simulation tool (developed by the National Renewable Energy Laboratory) with the load cases stimulated at under-rated, rated and above rated speeds. Then, the bivariate correction method is applied to model FWT extreme response for a 5-years return period prediction with a 95% confidence interval (CI), based on just 2 min short response record. The proposed methodology permits accurate correction of the bivariate correction of the proposed novel method's performance, it is concluded that the bivariate correction method can offer better robust and precise bivariate predictions of coupled surge motion and anchor tension of the FWT.

Keywords: floating wind turbine, FAST, bivariate correction method, extreme responses, bivariate probability distribution

1 INTRODUCTION

The environmental pollution caused by the massive use of fossil fuels has triggered the transformation of non-renewable energy to renewable energy. In 2020, global energy-related carbon dioxide emissions dropped by 5.8%, which reached the highest record of the percentage decrease since the second World War. Meanwhile, renewable energy reached 29%, the highest record share in global electricity (Secretariat, 2021). As one of the most promising renewable energy, wind energy has been evolving rapidly. 93 GW of new wind power installations was seen in 2020, which brought the global cumulative wind power capacity up to 743 GW (Council, 2021).

Although today's wind generation is mainly onshore, a great prospect is predicted for floating offshore wind. This is because floating wind turbines (FWTs) could harvest vast wind resources over deep water, where the wind capacity is more than four times the bottom-fixed wind. However, the high cost is the main challenge with the floating wind turbines. This is mainly due to the small scale of the floating wind farms and the

immaturity of the technology and supply chain. Significant technology improvement is a must to overcome the challenges and lower the levelized cost of energy (LCOE) of floating wind. Enhancing the insights into the loads and load effects of the FWTs will be an efficient solution to the successful deployment of floating wind. This is because a good understanding of the FWT dynamics can aid to reduce the risk and uncertainty associated with the wind turbine design and operating, thus reducing the turbine, foundation and operating expenses.

Mooring equipment is a crucial cost driver within the floating wind, while significant efforts need to be taken to develop a target safety level to realise a cost-effective design of mooring systems. Ultimate limit state (ULS) check serves as the main criteria for the mooring system design, which is performed based on extreme load effect analysis. The FWT is a complex combination with aerodynamics, hydrodynamics, control system and flexible structures and is subjected to various environmental conditions with wind, wave and currents. Determining the extreme values of the mooring system requires effective methods to deal with the extreme analysis problem.

Hsu et al. (Hsu et al., 2015) presented a comparative analysis in predicting the extreme mooring line tensions of an FWT exposed to a 100-years storm condition between model test and numerical simulations. The importance of the snap events to the mooring line tensions was investigated based on the exceedance probability analysis, and results showed that normalised snap loads tend to a constant 3, which can contribute to determining the safety factor of mooring lines. In the study of Hsu et al. (Hsu et al., 2017), a composite Weibull probability distribution was proposed for the mooring line dynamic tensions due to the effects of snap loads. The results showed that when the probability of the snap events is high, the developed composite Weibull distribution method could be effectively used to predict extreme tensions of FWT mooring systems. Xu et al. (Xu et al., 2019a) studied the influence of non-linear wave kinematics on fatigue damage and extreme responses of a 5MW semi-submersible FWT. Gumbel fitting and Average Conditional Exceedance Rate (ACER) methods were used to predict the extreme values of linear mooring tensions, and the results showed that the fully non-linear wave theory would lead to higher mooring line tensions than the linear wave theory. Cao et al. (Cao et al., 2020) investigated the extreme responses of a new concept of 10-MW semi-submersible FWT based on the experimental tests. The mean up-crossing rate method was used to estimate the extreme short-term values of the mooring linear tensions. The results showed that the extreme mooring line tensions are more likely to occur under harsh environmental conditions than operational conditions. Most of the previous studies predicting the extreme loads of mooring line tensions are based on the extreme value distribution models.

However, the extreme values predicted by this method depend on the tail of the probability distributions, which are pretty sensitive to uncertainties. The well developed ACER method can enhance the effectiveness and reliability in predicting the extreme values, but significant numerical or experimental efforts are needed to produce sufficient data for the ACER analysis. Motivated by this, it is essential to improve the method to predict extreme responses without devoting significant costs and efforts accurately.

This paper proposes a novel bivariate correction approach if the sensor malfunctions or the available data record is too short. The proposed method can more efficiently and reliably predict extreme



FIGURE 1 | The 10-MW OO-Star floating wind turbine (Yu et al., 2017).

loads in a 10 MW large floating wind turbine (FWT). More efficient and reliable estimations of extreme responses will better help predict the effects these loads have on the components allowing the development and implementation of a better design or control system for the FWT. Optimal wind turbine parameters would minimise potential FWT mechanical damage due to excessive environmental loadings (Xu et al., 2019a). Accurately predicted extreme loads will also allow the components to be more optimally sized. It contributes to more refined designs and lower failure rates, which is particularly important for the offshore wind industry as it advances the design, manufacturing and deployment of large FWTs (>10 MW) in the coming decade.

2 SYSTEM DESCRIPTION

A 10-MW FWT system (Yu et al., 2017) is used in this work, and is illustrated in **Figure 1**. The FWT system will be expounded in two parts in the following sections. Firstly, the reference wind turbine will be described, then the properties of the semi-submersible floater and the mooring system will be introduced **Figures 2** and **3**.

2.1 DTU 10-MW Reference Wind Turbine

The DTU 10-MW reference wind turbine (RWT) is used in this paper, designed from the NREL 5-MW RWT. The wind turbine was designed per the International Electrotechnical Commission (IEC) Class 1A wind regime and is a traditional three-bladed, clockwise rotation-upwind turbine equipping with a variable speed and collective pitch control system. The DTU 10-MW RWT numerical model has been





successfully developed and studied in many academic works, e.g., (Hu et al., 2021; Muggiasca et al., 2021; Wang et al., 2022; Yu et al., 2022). The summary of the DTU 10-MW RWT is shown in **Table 1**.

2.2 OO-Star Semi-Submersible Wind Floater and Mooring System

This work uses a semi-submersible floating structure to support the 10-MW RWT. It was introduced by Dr. techn. Olav Olsen

AS in the LIFES 50 + project (Yu et al., 2017). The floater comprises post-tensioned concrete, hosting a central column with three outer columns. The four columns are mounted on a star-shaped pontoon, where a slab is attached at the bottom. Three catenary mooring lines are used to maintain the floater in position, and in each line, a clumped mass is attached, separating the line into two segments. Greater details of the OO-Star Wind Floater and the mooring system are shown in **Table 2** and **Table 3**, respectively.

Parameter	Value		
Rating	10-MW		
Туре	Upwind/3 blades		
Control	Variable speed, collective pitch		
Drivetrain	Medium-speed, multiple stage		
	gearbox		
Cut-in, rated and cut-out wind speed (m/s)	4, 11.4, 25		
Minimum and maximum rotor speed (rpm)	6.0, 9.6		
Maximum generator speed (rpm)	480		
Rotor diameter (m)	178.3		
Hub height (m)	119.0		
Rotor mass (kg)	227962		
Nacelle mass (kg)	446036		
Tower mass (kg)	1.257×10^{6}		

TABLE 2 | The main properties for the 10-MW OO-Star wind floater.

Value
130
22
11
24158
21709
1.4462 × 10 ¹⁰
1.63 × 10 ¹⁰
15.23
14.236

TABLE 3 | The main properties for the mooring system of the 10-MW FWT.

Parameter	Value	
Radius to anchors from platform centerline (m)	691	
Anchor position below MSL (m)	130	
Initial vertical position of clump mass below MSL (m)	90.45	
Initial radius to clump mass from centerline (m)	148.6	
Length of clump mass upper segment (kg)	118	
Length of clump mass lower segment (kg)	585	
Equivalent weight per length in water (N/m)	3,200.6	
Extensional stiffness (N/m)	1.506 × 10 ⁹	

3 METHODOLOGY

This section describes the methodology adopted by authors to address engineering challenges related to safe and reliable design of FWTs (floating wind turbines). Note that the proposed ACER (Average Conditional Exceedance Rate) method along with The FAST simulation tool was already recently successfully used by the authors, see, e.g., (Xu et al., 2020).

3.1z Aero-Hydro-Elastic-Servo Dynamic Analysis of the 10-MW FWT

FAST (Fatigue, Aerodynamics, Structures and Turbulence) (version8, v8.16.00a-bjj), an open-source WT simulation tool developed by the National Renewable Energy Laboratory

(NREL), is utilised in this work for the fully coupled aerohydro-elastic-servo dynamic analysis for the 10-MW FWT. The FAST code couples together five computer codes: AeroDyn (Moriarty and Hansen, 2005), HydroDyn (Jonkman et al., 2014), ServoDyn, and MoorDyn (Hall, 2015), to account for the aerodynamic loads on rotor blades, hydrodynamic loads on floaters, control dynamics, structural dynamics and mooring system dynamics. In addition, FAST provides the interface for reading the time-varying stochastic wind for time-domain simulations. The FAST simulation tool has been successfully used in other well-known projects such as OC3: Offshore Code Comparison Collaboration (Jonkman and Musial, 2010) and OC4: IEA Task Wind 30 (Robertson et al., 2014), and its modelling capability has been authenticated using multiple floating structures in the Netherlands (Coulling et al., 2013).

3.1.1 Aerodynamics

The blades aerodynamic loads are calculated based on the quasisteady Blade Element Momentum (BEM) theory. BEM theory combines momentum theory and blade element theory. Various advanced corrections, including tip loss, hub loss, skewed inflow and dynamic stall corrections, are included in the BEM method. The Prandtl corrections are implemented to account for the hub and blade tip losses due to a finite number of blades. The Glauert correction is applied to account for the induction factors, while the Pitt and Peters' model accounts for the skewed inflow correction. The dynamic stall correction is employed in the Beddoes-Leishman model. More details about the aerodynamic load calculation in the FAST code can be seen in the AeroDyn theory manual (Moriarty and Hansen, 2005).

3.1.2 Hydrodynamics

Hydrodynamic loads acting on the semi-submersible floater are calculated based on potential flow theory with Morison's drag term considered. It accounts for the wave pressures and viscous loads, respectively. Hydrodynamic coefficients, such as added mass and potential damping coefficients, and first-order wave excitation load transfer function are firstly estimated in the frequency domain by a panel code, WAMIT, according to the potential flow theory. These hydrodynamic coefficients are then transformed into the time domain using the convolution technique.

3.1.3 Structural Dynamics

A combined multi-body and modal structural approach is considered in the FAST code to account for the structural dynamics of the FWT. The blades, tower and driveshaft are considered flexible bodies, while the nacelle, hub and floater are rigid bodies. The inherent structural damping in the blades and tower are represented using the Rayleigh damping model. The structural dynamic responses in the time-domain are calculated by solving the equations of motion of the rigidflexible coupled system derived by Kane's approach, see (Kane and Levinson, 1983).

3.1.4 Control System Dynamics

The control system used in the 10-MW FWT is implemented in two operational modes: the below-rated and full-rated

regions. The generator torque-speed curve regulates the rotor rotational speed with an optimal tip speed ratio in the belowrated region, achieving maximum power generation. A proportional-integral (PI) algorithm regulates the blade pitch angle to reduce the structural loading while keeping the rated power generation in the full-rated region. The PI parameters are modified from the land-based RWT to avoid the negative damping effects, which are essential in affecting the platform motions for FWTs.

3.2 ACER Method

Consider a long term global response process (such as anchor tension force or FWT surge motion) X (t) of the FWT, measured over a time interval (0, T). Let X_1, \ldots, X_N be measurements of the process X(t) at discrete time moments t_1, \ldots, t_N in (0, T). The target is to estimate the distribution function of the extreme value $M_N = \max \{X_j; j = 1, \ldots, N\}$ accurately namely, to estimate CDF (cumulative density function) $P(\eta) =$ Prob $(M_N \le \eta)$ for large values of the response η . The following random functions are introduced:

$$A_{kj}(\eta) = \mathbf{1} \{ X_j > \eta, X_{j-1} \le \eta, \dots, X_{j-k+1} \le \eta \}, \quad j = k, \dots, N, k$$

= 2, 3,...

and

$$B_{kj}(\eta) = \mathbf{1} \{ X_{j-1} \le \eta, \dots, X_{j-k+1} \le \eta \}, \quad j = k, \dots, N, \ k$$

= 2, 3, \dots, (2)

Where 1{A} = 1 if A is true, while it is 0 if not. As shown in (Wang, 2001; Bak et al., 2013; Naess and Karpa, 2015a; Naess and Karpa, 2015b; Gaidai et al., 2016; Jian et al., 2018; Gaidai et al., 2019a; Gaidai et al., 2019b; Xu et al., 2019b; Gaidai et al., 2020; Xu et al., 2021)

$$P_{k}(\eta) \approx \exp \left(-\sum_{j=k}^{N} \frac{\mathbb{E}\left[A_{kj}(\eta)\right]}{\mathbb{E}\left[B_{kj}(\eta)\right]}\right) \approx \exp \left(-\sum_{j=k}^{N} \mathbb{E}\left[A_{kj}(\eta)\right]\right), \quad \eta \to \infty$$
(3)

The measured time series can be subdivided into *K* subsequent (short term) blocks such that $E[A_{kj}(\eta)]$ remains approximately constant within each block and such that $\sum_{j \in C_i} E[A_{kj}(\eta)] \approx \sum_{j \in C_i} a_{kj}(\eta)$ for a sufficient range of η -values, where C_i stands for the set of indices for the block with number *i*; with i = 1, ..., K, and where $a_{kj}(\eta)$ are the realised values of $A_{kj}(\eta)$ for the measured time series, then $\sum_{j=k}^{N} E[A_{kj}(\eta)] \approx \sum_{j=k}^{N} a_{kj}(\eta)$. Thus, for a given stationary process (short term sea current state), one has

$$P_k(\eta) \approx \exp\left(-(N-k+1)\hat{\varepsilon}_k(\eta)\right) \tag{4}$$

Where

$$\hat{\varepsilon}_k(\eta) = \frac{1}{N-k+1} \sum_{j=k}^N a_{kj}(\eta) \tag{5}$$

In the above equations, an assumption of ergodicity has been used for each short term segment of the recorded time series in order to estimate the short term expected values by using observed values of the $a_{kj}(\eta)$ functions. An alternative way of expressing the long term extreme value distribution in **Eqn. 4**, is obtained by considering the empirical probability distribution of m = 1, ..., M sea current states having probabilities p_m , so that $\sum_{m=1}^{M} p_m = 1$. Next, introduce the long term ACER function of order k:

$$ACER_{k}(\eta) \equiv \sum_{m=1}^{M} \hat{\varepsilon}_{k}(\eta, m) p_{m}$$
(6)

Where $\hat{\varepsilon}_k(\eta, m)$ is the same function as in **Eqn. 5** but restricted to a specific sea state with number *m*. As shown in (Gaidai et al., 2020)- (Xu et al., 2019b), the long-term extreme value distribution of M(T), can then be expressed as follows based on the ACER function of order *k*:

$$P(\eta) \approx \exp(-N.\mathrm{ACER}_k(\eta))$$
 (7)

Where $ACER_k(\eta)$ is the long term empirical ACER function of order *k*, with *k* "*N*; *N* is the total number of data points from the recorded time series used to estimate the ACER functions. Typically, these could be local peaks from the measured time series.

As the order *k* increases, the accuracy of **Eqn.** 7 improves and $ACER_k(\eta)$ functions converge conveniently fast with growing k, see (Gaidai et al., 2020)- (Xu et al., 2019b). Note that by increasing the conditioning level *k*, possible data clustering effects, e.g. with narrow-band response components in the recorded time series, can be accounted for. This is an essential advantage of the ACER method as it increases the accuracy of its extreme predictions and avoids resulting in over-conservative design values.

The $ACER_k$ as functions of the level, η are generally quite regular in the tail, i.e., for high response values of η . More specifically, for $\eta \ge \eta_0$, the tail behaves very closely like $exp\{-(a\eta + b)^c + d\}$ with *a*, *b*, *c*, *d* being suitable constants.

It is suggested to do the optimisation on the log-level by minimising the following mean square error function F with respect to the four arguments a_k , b_k , c_k , p_k , q_k .

$$F(a_k, b_k, c_k, p_k, q_k) = \int_{\eta_0}^{\eta_1} \omega(\eta) \{ \ln(ACER_k(\eta)) - d_k + (a_k\eta + b_k)^{c_k} \}^2 d\eta, \quad \eta \ge \eta_0$$
(8)

Where η_1 is a suitable data cut-off value, i.e., the largest η response value with the confidence with acceptable interval width. Weight function ω is defined as $\omega(\eta) = \{lnC^+(\eta) - lnC^-(\eta)\}^{-2}$ with $(C^-(\eta), C^+(\eta))$ being a 95% CI, empirically estimated from the measured data. Detailed procedure for further optimising parameters a_k , b_k , c_k , d_k has been outlined in (Gaidai et al., 2020)- (Xu et al., 2019b).

(1)

3.3 Bivariate Correction

As already mentioned, the method proposed in this paper will be based on the ACER methodology. This involves both the univariate ACER functions and the bivariate ACER functions. The unique feature of the ACER functions is that they provide the possibility to portray the exact extreme value distribution inherent in the data time series, both the univariate and the bivariate (Wang, 2001)- (Xu et al., 2019b). Hence, the ACER method is fundamentally different from the traditional approach relying on the fitting of hardly asymptotic data to asymptotic extreme value distributions, which are based on the assumption of stationary time series instead of the ACER method. The empirical ACER functions are represented in nonparametric numerical functions with uncertainty bounds. The accuracy obtained depends, of course, on the amount of data available to estimate these functions. It is also an essential feature of the ACER method that it is not limited to stationary time series. It is entirely valid for nonstationary time series as long as the measured data reflects this nonstationarity. For the sake of easy reference, the univariate ACER methodology has been briefly outlined below. The bivariate case follows quite closely the univariate case.

This section presents a statistical bivariate integral correction based on the bivariate ACER method coupled with the Gumbel logistic model (Gaidai et al., 2019a)- (Xu et al., 2019b). Note that this correction is not limited to only extreme value estimates. However, it can be applied with appropriate bivariate models for any statistical values of interest to improve their accuracy based on synchronously measured longer, highly correlated data records.

Let $X = \max\{R_2(t); t \in [0, T_{\text{return}}]\}$, $Y = \max\{R_3(t); t \in [0, T]\}$, where X = Surge and Y = Tension are responses simulated synchronously, see **Section 3.1**, with *T* being return period of interest, and let $F_{XY}(\xi, \eta) = Prob(X \le \xi, Y \le \eta)$ be the joint bivariate cumulative distribution function (CDF) of (X, Y).

Raw response time series with dt = 0.0246 sec were blocked into k = 10 consecutive discrete time points maxima to reduce the neighbouring data points correlation effect. The conditioning level k was chosen according to the response power spectral density as it was observed that ACER functions have converged at that level in the distribution tail, see Section 3.2.

 $F_X(\xi)$ and $F_Y(\eta)$ denote the corresponding univariate marginal CDFs for X and Y, respectively. In this paper, it is assumed that the bivariate couple (Surge(t), Tension(t)) has been observed over a period of time $t \in [0, \tilde{T}]$, where the observation duration \tilde{T} is not long enough for accurately predicting the univariate extreme response levels with a target low probability of interest. Now, consider the case when a 'long' record of Tension(t) is available over a time $t \in [0, T]$, with $T''\tilde{T}$, with a corresponding estimated CDF $F_X^{long}(\xi)$ of the CDF $F_X(\xi)$, which has a probability density function (PDF) $p_X = F'_X$. In this paper $T_{\text{return}} \approx 5$ years is the return period that corresponds to the extreme probability p of interest, $p = 0.5\Delta 10^{-7}$, T = 20 h, $\tilde{T} = T/600$ see **Section 4**. Then for any Y- response level of interest η_* , with $\Delta \to 0$,

$$F_{Y}(\eta_{*}) = F_{XY}\left(E,\eta\right) = \int_{0}^{+\infty} \operatorname{Prob}\left(Y \le \eta_{*} \middle| X = \xi\right) p_{X}(\xi) d\xi$$
$$= \int_{0}^{+\infty} \frac{\operatorname{Prob}\left(Y \le \eta_{*}, X \in [\xi, \xi + \Delta]\right)}{\operatorname{Prob}\left(X \in [\xi, \xi + \Delta]\right)} p_{X}(\xi) d\xi$$
$$= \int_{0}^{+\infty} F'_{XY,X}(\xi,\eta_{*}) d\xi, \tag{9}$$

with $F'_{XY,X}(\xi,\eta) = \frac{\partial}{\partial\xi}F_{X,Y}(\xi,\eta)$. See (Gaidai et al., 2019a)- (Xu et al., 2019b) for the details related to these equations. The following copula model for the bivariate extreme value distribution is referred to as the Gumbel logistic model (Wang, 2001)- (Xu et al., 2019b):

$$F_{XY}(\xi,\eta) = \exp\{-\left[(-\ln F_X(\xi))^m + (-\ln F_Y(\eta))^m\right]^{1/m}\}$$
(10)

In this model, it is seen that m = 1 corresponds to the case when X and Y are independent. When 0 < m < 1, X and Y become dependent, (Wang, 2001)- (Xu et al., 2019b). However, this dependence structure is of a special kind since it only involves the marginal distributions. Still, it appears to be useful in some practical cases. Another popular extreme value copula is the Asymmetric Gumbel logistic model, see (Wang, 2001)- (Xu et al., 2019b). This bivariate extreme value model will not be discussed further in this paper since it gives identical results as the Gumbel logistic model. The Gumbel logistic model has been verified to be useful for various offshore engineering practical applications, provided the marginal extreme value distributions are estimated using the univariate ACER method instead of standard asymptotic extreme value distributions (Wang, 2001)- (Xu et al., 2019b). This is because the asymptotic distributions typically used are often not accurate enough in the tails, being the results of fitting real, subasymptotic data to asymptotic distributions. If Eqn. 35 is differentiated with respect to x, it is obtained that

$$F'_{XY,X}(\xi,\eta_*) = F_{XY}(\xi,\eta_*) [1 + (\ln F_Y(\eta_*) / \ln F_X(\xi))^m]^{\frac{1}{m}} \frac{1}{d\xi} \ln F_X(\xi)$$
(11)

The numerical estimates $\hat{F}_Y(\eta_*)$ and $\hat{F}_{XY}(\xi, \eta_*)$ of $F_Y(\eta_*)$ and $F_{XY}(\xi, \eta_*)$, respectively, based on the available time series of recorded data, are now used in the following expression to obtain the corrected estimate $F_Y^{corr}(\eta_*)$:

$$F_{Y}^{corr}\left(\eta_{*}\right) = \int_{0}^{+\infty} \hat{F}_{XY}\left(\xi,\eta_{*}\right) \left[1 + \left(\ln \hat{F}_{Y}\left(\eta_{*}\right) / \ln \hat{F}_{X}^{long}\left(\xi\right)\right)^{m}\right]^{\frac{1}{m}-1} \frac{d}{d\xi} \ln \hat{F}_{X}^{long}\left(\xi\right) d\xi.$$
(12)

Note that all quantities on the right side of **Eqn. 12** are known from the available time series of recorded data. The Gumbel copula parameter *m* has been calibrated to fit joint empirical distribution \hat{F}_{XY} . For the latter optimisation task, the Trustregion-reflective non-linear least-squares optimisation algorithm can be used, and the interior-point algorithm to find the minimum of constrained non-linear multivariable function; for details, see (Wang, 2001)- (Xu et al., 2019b).





FIGURE 5 | ACER functions for shorter and longer FWT response records: Surge and Tension FWT responses, upper and lower figures, respectively. The ACER function decimal logarithm values are on the vertical axes. Mean current velocity 8 m/s.

TABLE 4 Correction results for measured FWT res	ponse: 20 h (long) and 2 min (short).	. Five years return period prediction.
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	Mean wind 8 m/s	Mean wind 12 m/s	Mean wind 16 m/s
$T_{\rm short}/T_{\rm long}$	0.72	0.91	0.51
$T_{\rm corrected}/T_{\rm long}$	0.97	1.30	0.82

4 BIVARIATE CORRECTION RESULTS

Three realistic mean wind velocities of 8, 12 and 16 m/s are studied in this paper. For the three environmental conditions, 20 different random samples of wind and wave are applied for each sea state. Each simulation lasts 4000s, where the first 400s is removed to reduce the transient effect induced by the wind turbine start-up. Therefore, 1-h data in each simulation is formed and is used for extreme value analysis in this work. The results shown in this work are based on the average of 20 1-h simulations to reduce the stochastic variability.

Simulated surge motion and anchor tension, *Surge* and *Tension* correspondingly, FWT responses possess a high correlation coefficient $R_{corr} \approx 0.9$. This paper studies synchronous measurements of Surge and Tension FWT responses, also referred to in figures as X and Y, respectively. It is seen from **Figure 4** that there is a non-linear dependency between *Surge* and *Tension* FWT responses, therefore task of bivariate prediction is not trivial.

Figure 5 shows the ACER functions for the more extended observation period *T*, which is 20 h of numerical simulation, and for the shorter period $\tilde{T} = T/600$. It is seen that, due to the high correlation between the two FWT response processes, both overestimate the ACER function levels compared with the more extended dataset ACER curve.

The following results were obtained for the simulated FWT response motions bivariate correction.

Table 4 presents the correction results for the corrected FWT anchor *Tension*. There are no monotonic relationship between the results and mean wind speeds. It is seen that the proposed correction technique resulted in remarkable improvement in accuracy, from about 90% over-prediction down to 30% in the case of mean wind velocity 12 m/sec. The return period for the predicted response level was chosen to be 5 years. The latter provides a practical example, supporting the novel correction technique introduced in this paper.

5 CONCLUSION

The FWT surge motion and anchor tension force due to environmental wind and wave loads were studied for three operating conditions of mean wind speeds of 8, 12 and 16 m/s.

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The bivariate correction method was briefly described and applied to account for the coupled load effects, namely surge motion and anchor tension force simulated synchronously in time.

This paper proposed using the bivariate correction method to investigate the extreme structural responses (motion surge and anchor tension force) during FWT realistic operation. High correlation between the two processes is the key requirement for the described correction to achieve improved prediction accuracy. As shown in the presented study, there is a practical advantage in applying the bivariate correction introduced in this paper, as it brings prediction based on short time series of data quite close to the prediction based on a much longer time series. Thus, a significant improvement in extreme value prediction accuracy is obtained. Some practical situations that may justify the above-mentioned analysis would be:

- Another similar FWT is being designed for the same environmental condition, then data collected from one FWT may be useful for another.
- Malfunctioning of one measuring sensor, while another is well-functioning.

This paper shows that applying the bivariate correction for the particular cases studied has increased extreme FWT response prediction accuracy. This improvement shows that the proposed correction method can be useful in engineering design in case of sensor malfunctioning or available data record is too short.

DATA AVAILABILITY STATEMENT

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

AUTHOR CONTRIBUTIONS

OG, Writing, Conceptualisation, Method, Analysis, Discussion. YX, Writing, Conceptualisation, Method, Analysis. FW, Funding, Review, Discussions. SW, Analysis, Writing. PY, Discussions. AN, Review, Discussions.

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